



# Workshop 8.2a: Heterogeneity

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## 1. Linear modelling assumptions

### 1.1. Assumptions

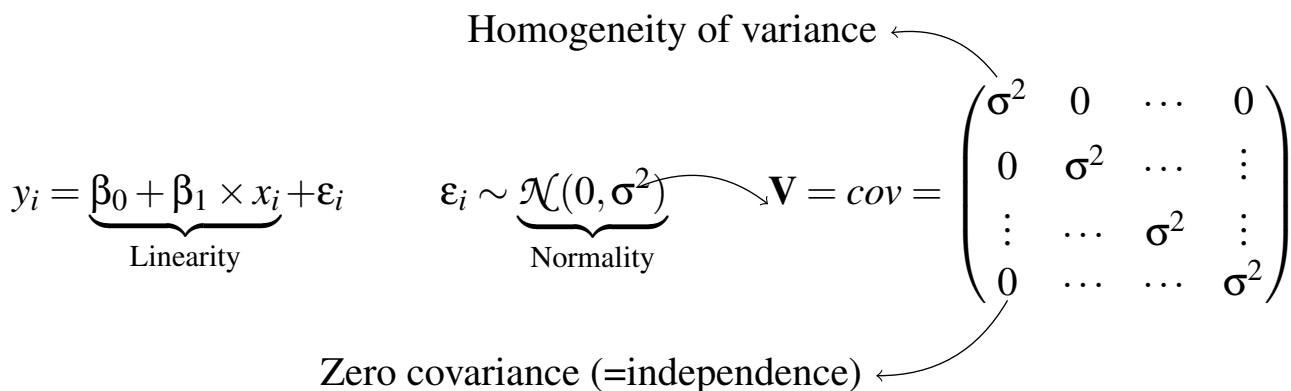
$$y_i = \beta_0 + \beta_1 \times x_i + \epsilon_i$$

$$\epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

### 1.2. Linear modelling assumptions

$$y_i = \beta_0 + \beta_1 \times x_i + \epsilon_i$$

$$\epsilon_i \sim \mathcal{N}(0, \sigma^2)$$



### 1.3. Dealing with Heterogeneity

y	x
41.9	1
48.5	2
43	3
51.4	4



y	x
51.2	5
37.7	6
50.7	7
65.1	8
51.7	9
38.9	10
70.6	11
51.4	12
62.7	13
34.9	14
95.3	15
63.9	16

## 1.4. Dealing with Heterogeneity

```
> data1 <- read.csv('..../data/D1.csv')
> summary(data1)
```

y	x
Min. :34.90	Min. : 1.00
1st Qu.:42.73	1st Qu.: 4.75
Median :51.30	Median : 8.50
Mean :53.68	Mean : 8.50
3rd Qu.:63.00	3rd Qu.:12.25
Max. :95.30	Max. :16.00

$$y_i = \beta_0 + \beta_1 \times x_i + \epsilon_i \\ \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

- estimate  $\beta_0$ ,  $\beta_1$  and  $\sigma^2$

## 1.5. Dealing with Heterogeneity

## 1.6. Dealing with Heterogeneity



## *1.7. Dealing with Heterogeneity*

## *1.8. Dealing with Heterogeneity*

$$y_i = \underbrace{\beta_0 + \beta_1 \times x_i}_{\text{Linearity}} + \epsilon_i$$

## Homogeneity of variance

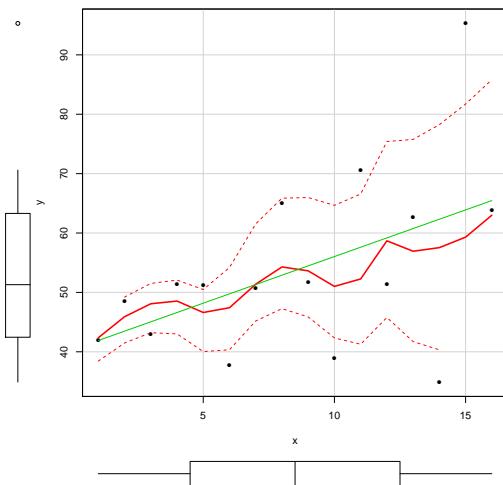
$$\varepsilon_i \sim \underbrace{\mathcal{N}(0, \sigma^2)}_{\text{Normality}}$$

$$\begin{pmatrix} \sigma^2 & 0 & \cdots & 0 \\ 0 & \sigma^2 & \cdots & \vdots \\ \vdots & \cdots & \sigma^2 & \vdots \\ 0 & \cdots & \cdots & \sigma^2 \end{pmatrix}$$

Zero covariance (=independence)

$$\mathbf{V} = \sigma^2 \times \underbrace{\begin{pmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & \vdots \\ \vdots & \cdots & 1 & \vdots \\ 0 & \cdots & \cdots & 1 \end{pmatrix}}_{\text{Identity matrix}} = \underbrace{\begin{pmatrix} \sigma^2 & 0 & \cdots & 0 \\ 0 & \sigma^2 & \cdots & \vdots \\ \vdots & \cdots & \sigma^2 & \vdots \\ 0 & \cdots & \cdots & \sigma^2 \end{pmatrix}}_{\text{Variance-covariance matrix}}$$

### 1.9. Dealing with Heterogeneity



- variance proportional to X
- variance inversely proportional to X

### 1.10. Dealing with Heterogeneity

- variance inversely proportional to X

$$\mathbf{V} = \sigma^2 \times X \times \underbrace{\begin{pmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & \vdots \\ \vdots & \cdots & 1 & \vdots \\ 0 & \cdots & \cdots & 1 \end{pmatrix}}_{\text{Identity matrix}} = \underbrace{\begin{pmatrix} \sigma^2 \times \frac{1}{\sqrt{X_1}} & 0 & \cdots & 0 \\ 0 & \sigma^2 \times \frac{1}{\sqrt{X_2}} & \cdots & \vdots \\ \vdots & \cdots & \sigma^2 \times \frac{1}{\sqrt{X_i}} & \vdots \\ 0 & \cdots & \cdots & \sigma^2 \times \frac{1}{\sqrt{X_n}} \end{pmatrix}}_{\text{Variance-covariance matrix}}$$

### 1.11. Dealing with Heterogeneity

$$\mathbf{V} = \sigma^2 \times \omega, \quad \text{where } \omega = \underbrace{\begin{pmatrix} \frac{1}{\sqrt{X_1}} & 0 & \cdots & 0 \\ 0 & \frac{1}{\sqrt{X_2}} & \cdots & \vdots \\ \vdots & \cdots & \frac{1}{\sqrt{X_i}} & \vdots \\ 0 & \cdots & \cdots & \frac{1}{\sqrt{X_n}} \end{pmatrix}}_{\text{Weights matrix}}$$



## 1.12. Dealing with Heterogeneity

Calculating weights

```
> 1/sqrt(data1$x)
```

```
[1] 1.0000000 0.7071068 0.5773503 0.5000000 0.4472136 0.4082483 0.3779645 0.3535534 0.3333333
[10] 0.3162278 0.3015113 0.2886751 0.2773501 0.2672612 0.2581989 0.2500000
```

## 1.13. Generalized least squares (GLS)

1. use OLS to estimate fixed effects
2. use these estimates to estimate variances via ML
3. use these to re-estimate fixed effects (OLS)

## 1.14. Generalized least squares (GLS)

ML is biased (for variance) when N is small:

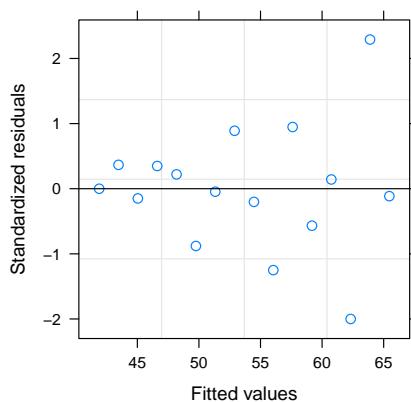
- use REML
- max. likelihood of residuals rather than data

## 1.15. Variance structures

Variance function	Variance structure	Description
<code>varFixed(~ x)</code>	$V = \sigma^2 \times x$	variance proportional to 'x' (the covariate)
<code>varExp(form= ~ x)</code>	$V = \sigma^2 \times e^{2\delta \times x}$	variance proportional to the exponential of 'x' raised to a constant power
<code>varPower(form= ~ x)</code>	$V = \sigma^2 \times  x ^{2\delta}$	variance proportional to the absolute value of 'x' raised to a constant power
<code>varConstPower(form= ~ x)</code>	$V = \sigma^2 \times (\delta_1 +  x ^{\delta_2})^2$	a variant on the power function
<code>varIdent(form= ~ A)</code>	$V = \sigma^2 \times I$	when A is a factor, variance is allowed to be different for each level ( $j$ ) of the factor
<code>varComb(form= ~ x   A)</code>	$V = \sigma^2 \times x \times I$	combination of two of the above

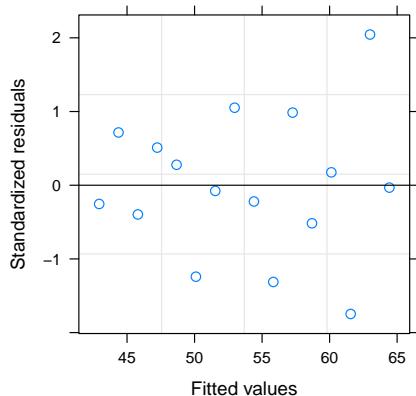
## 1.16. Generalized least squares (GLS)

```
> library(nlme)
> data1.gls <- gls(y~x, data1,
+                      method='REML')
> plot(data1.gls)
```



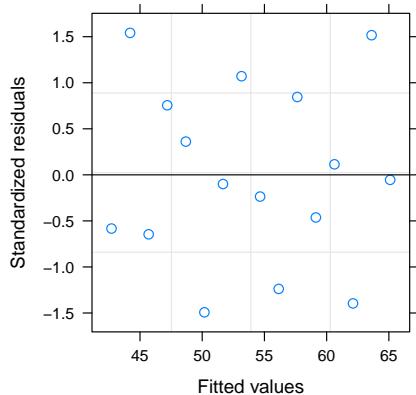


```
> library(nlme)
> data1.gls1 <- gls(y~x, data=data1, weights=varFixed(~x),
+                      method='REML')
> plot(data1.gls1)
```



## 1.17. Generalized least squares (GLS)

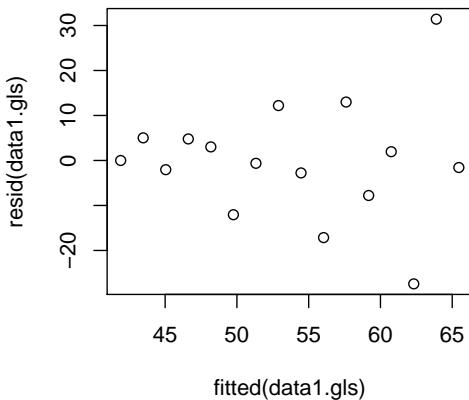
```
> library(nlme)
> data1.gls2 <- gls(y~x, data=data1, weights=varFixed(~x^2),
+                      method='REML')
> plot(data1.gls2)
```



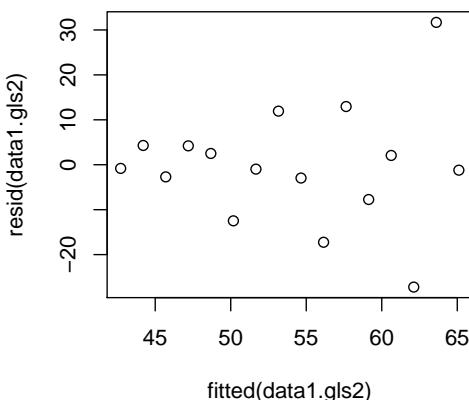
## 1.18. Generalized least squares (GLS)

### 1.18.1. wrong

```
> plot(resid(data1.gls) ~
+       fitted(data1.gls))
```



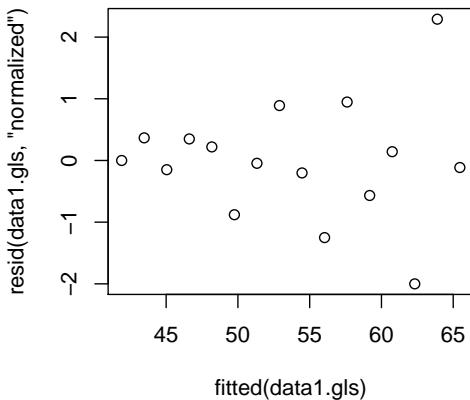
```
> plot(resid(data1.gls2) ~  
+       fitted(data1.gls2))
```



## 1.19. Generalized least squares (GLS)

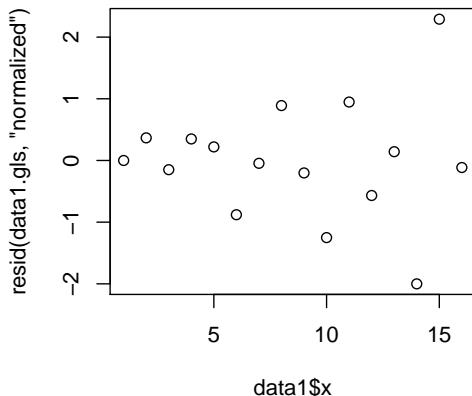
### 1.19.1. CORRECT

```
> plot(resid(data1.gls,'normalized') ~  
+       fitted(data1.gls))
```

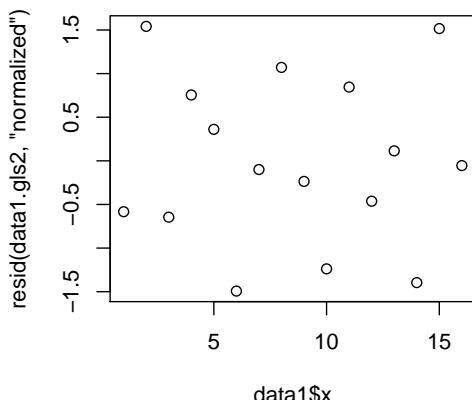




Australian Government



```
> plot(resid(data1.gls2,'normalized') ~ data1$x)
```



## 1.21. Generalized least squares (GLS)

```
> AIC(data1.gls, data1.gls1, data1.gls2)
```

	df	AIC
data1.gls	3	127.6388
data1.gls1	3	121.0828
data1.gls2	3	118.9904

```
> library(MuMin)
> AICc(data1.gls, data1.gls1, data1.gls2)
```

	df	AICc
data1.gls	3	129.6388
data1.gls1	3	123.0828
data1.gls2	3	120.9904

```
> #OR
> anova(data1.gls, data1.gls1, data1.gls2)
```

**Townsville address:** PMB No 3,  
Townsville MC, Qld 4810  
Tel: (07) 4753 4444  
Fax: (07) 4772 5852

**Darwin address:** PO Box No 41775,  
Casuarina NT 0811  
Tel: (08) 8920 9240  
Fax: (08) 8920 9222  
[www.aims.gov.au](http://www.aims.gov.au)

**Perth address:** The UWA Oceans Institute (M096)  
35 Stirling Highway, Crawley WA 6009  
Tel: (08) 6369 4000  
Fax: (08) 6488 4585



```
Model df      AIC      BIC  logLik
data1.gls     1 127.6388 129.5559 -60.81939
data1.gls1    2 121.0828 123.0000 -57.54142
data1.gls2    3 118.9904 120.9076 -56.49519
```

## 1.22. Generalized least squares (GLS)

```
> summary(data1.gls)
```

Generalized least squares fit by REML

```
Model: y ~ x
Data: data1
AIC      BIC  logLik
127.6388 129.5559 -60.81939
```

Coefficients:

	Value	Std. Error	t-value	p-value
(Intercept)	40.33000	7.189442	5.609615	0.0001
x	1.57074	0.743514	2.112582	0.0531

Correlation:

```
(Intr)
x -0.879
```

Standardized residuals:

Min	Q1	Med	Q3	Max
-2.00006105	-0.29319830	-0.02282621	0.35357567	2.29099872

Residual standard error: 13.70973

Degrees of freedom: 16 total; 14 residual

```
> summary(data1.gls2)
```

Generalized least squares fit by REML

```
Model: y ~ x
Data: data1
AIC      BIC  logLik
118.9904 120.9075 -56.49519
```

Variance function:

```
Structure: fixed weights
Formula: ~x^2
```

Coefficients:

	Value	Std. Error	t-value	p-value
(Intercept)	41.21920	1.493556	27.598018	0.0000
x	1.49282	0.469988	3.176287	0.0067

Correlation:

```
(Intr)
x -0.671
```

Standardized residuals:

Min	Q1	Med	Q3	Max
-----	----	-----	----	-----



-1.49259798 -0.59852829 -0.07669281 0.77799410 1.54157863

Residual standard error: 1.393108

Degrees of freedom: 16 total; 14 residual

### 1.23. Generalized least squares (GLS)

```
> data1$cx <- scale(data1$x, scale=FALSE)
> data1.gls <- gls(y~cx, data1,
+                      method='REML')
> summary(data1.gls)
```

Generalized least squares fit by REML

Model: y ~ cx

Data: data1

	AIC	BIC	logLik
cx	127.6388	129.5559	-60.81939

Coefficients:

	Value	Std. Error	t-value	p-value
(Intercept)	53.68125	3.427432	15.662236	0.0000
cx	1.57074	0.743514	2.112582	0.0531

Correlation:

	(Intr)
cx	0

Standardized residuals:

	Min	Q1	Med	Q3	Max
	-2.00006105	-0.29319830	-0.02282621	0.35357567	2.29099872

Residual standard error: 13.70973

Degrees of freedom: 16 total; 14 residual

```
> data1$cx <- scale(data1$x, scale=FALSE)
> data1.gls2 <- gls(y~cx, data1,
+                      weights=varFixed(~x^2), method='REML')
> summary(data1.gls2)
```

Generalized least squares fit by REML

Model: y ~ cx

Data: data1

	AIC	BIC	logLik
cx	118.9904	120.9075	-56.49519

Variance function:

Structure: fixed weights

Formula: ~x^2

Coefficients:

	Value	Std. Error	t-value	p-value
(Intercept)	53.90814	3.190165	16.898231	0.0000
cx	1.49282	0.469988	3.176287	0.0067

Correlation:

	(Intr)
--	--------



cx 0.938

Standardized residuals:

Min	Q1	Med	Q3	Max
-1.49259798	-0.59852829	-0.07669281	0.77799410	1.54157863

Residual standard error: 1.393108

Degrees of freedom: 16 total; 14 residual

## 2. Heteroscedadacy in ANOVA

### 2.1. Heteroscedadacy in ANOVA

```
> data2 <- read.csv('~/data/D2.csv')
> summary(data2)
```

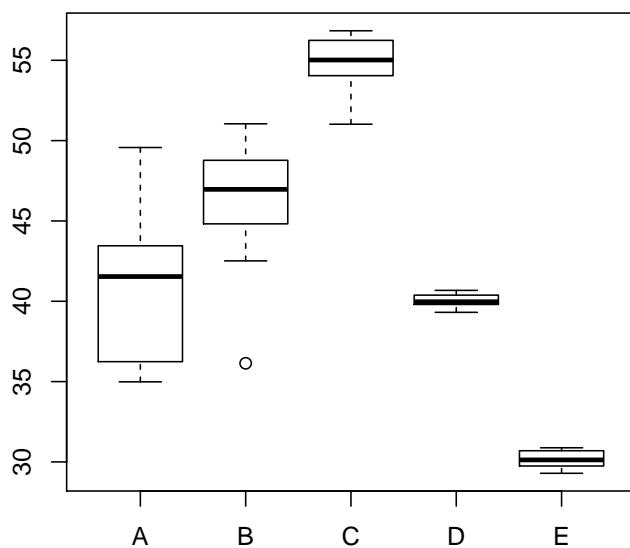
y	x
Min. :29.29	A:10
1st Qu.:36.17	B:10
Median :40.89	C:10
Mean :42.34	D:10
3rd Qu.:49.32	E:10
Max. :56.84	

$$\begin{aligned}y_i &= \mu + \alpha_i + \varepsilon_i \\ \varepsilon_i &\sim \mathcal{N}(0, \sigma^2)\end{aligned}$$

- estimate  $\mu$ ,  $\alpha_i$  and  $\sigma^2$

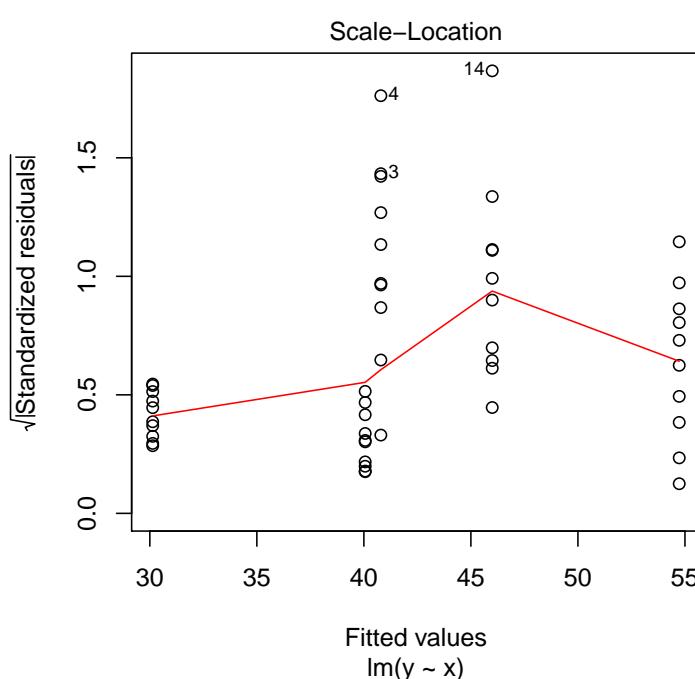
### 2.2. Heteroscedadacy in ANOVA

```
> boxplot(y~x, data2)
```



### 2.3. Heteroscedacity in ANOVA

```
> plot(lm(y~x, data2), which=3)
```





## 2.4. Heteroscedadacy in ANOVA

$$\varepsilon \sim \mathcal{N}(0, \sigma_i^2 \times \omega)$$

Effect ( $\alpha$ ) 1 ( $i=1$ )	$\begin{pmatrix} y_{1i} \\ y_{2i} \\ y_{3i} \end{pmatrix} = \begin{pmatrix} 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 0 & 0 \end{pmatrix} \times (\beta_i) + \begin{pmatrix} \varepsilon_{1i} \\ \varepsilon_{2i} \\ \varepsilon_{3i} \end{pmatrix} \quad \varepsilon_i \sim \mathcal{N}(0, \begin{pmatrix} \sigma_1^2 & 0 & 0 \\ 0 & \sigma_1^2 & 0 \\ 0 & 0 & \sigma_1^2 \end{pmatrix})$
Effect ( $\alpha$ ) 2 ( $i=2$ )	$\begin{pmatrix} y_{1i} \\ y_{2i} \\ y_{3i} \end{pmatrix} = \begin{pmatrix} 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 1 & 0 \end{pmatrix} \times (\beta_i) + \begin{pmatrix} \varepsilon_{1i} \\ \varepsilon_{2i} \\ \varepsilon_{3i} \end{pmatrix} \quad \varepsilon_i \sim \mathcal{N}(0, \begin{pmatrix} \sigma_2^2 & 0 & 0 \\ 0 & \sigma_2^2 & 0 \\ 0 & 0 & \sigma_2^2 \end{pmatrix})$
Effect ( $\alpha$ ) 3 ( $i=3$ )	$\begin{pmatrix} y_{1i} \\ y_{2i} \\ y_{3i} \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 & 1 \\ 1 & 0 & 0 & 1 \\ 1 & 0 & 0 & 1 \end{pmatrix} \times (\beta_i) + \begin{pmatrix} \varepsilon_{1i} \\ \varepsilon_{2i} \\ \varepsilon_{3i} \end{pmatrix} \quad \varepsilon_i \sim \mathcal{N}(0, \begin{pmatrix} \sigma_3^2 & 0 & 0 \\ 0 & \sigma_3^2 & 0 \\ 0 & 0 & \sigma_3^2 \end{pmatrix})$

## 2.5. Heteroscedadacy in ANOVA

$$V_{ji} = \begin{pmatrix} \sigma_1^2 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & \sigma_1^2 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & \sigma_1^2 & 0 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & \sigma_2^2 & 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & \sigma_2^2 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & \sigma_2^2 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & \sigma_2^2 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & \sigma_3^2 & 0 \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 & \sigma_3^2 \end{pmatrix}$$

## 2.6. Heteroscedadacy in ANOVA

```
> data2.sd <- with(data2, tapply(y,x, sd))
> 1/(data2.sd[1]/data2.sd)
```

A	B	C	D	E
1.00000000	0.91342905	0.40807277	0.08632027	0.12720488

## 2.7. Variance structures

Variance function	Variance structure	Description
<code>varFixed( x )</code>	$V = \sigma^2 \times x$	variance proportional to 'x' (the covariate)
<code>varExp(form= x)</code>	$V = \sigma^2 \times e^{2\delta \times x}$	variance proportional to the exponential of 'x' raised to a constant power
<code>varPower(form= x)</code>	$V = \sigma^2 \times  x ^{\delta}$	variance proportional to the absolute value of 'x' raised to a constant power
<code>varConstPower(form= x)</code>	$V = \sigma^2 \times (\delta_1 +  x ^{\delta_2})^2$	a variant on the power function
<code>varIdent(form=  A )</code>	$V = \sigma^2 \times I$	when A is a factor, variance is allowed to be different for each level ( $j$ ) of the factor
<code>varComb(form= x   A)</code>	$V = \sigma^2 \times x \times I$	combination of two of the above



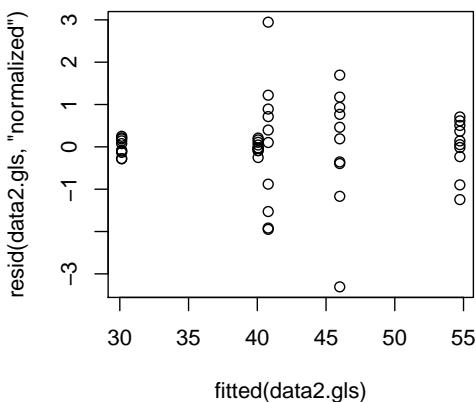
## 2.8. Heteroscedadacy in ANOVA

```
> library(nlme)
> data2.gls <- gls(y~x, data=data2,
+   method="REML")
```

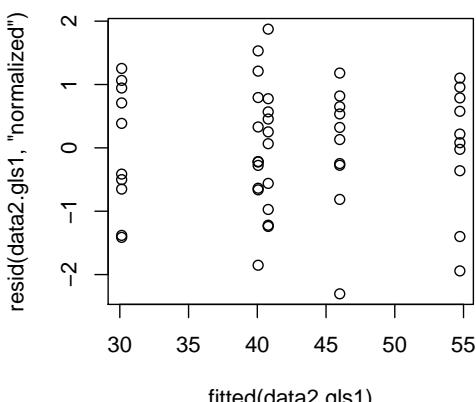
```
> library(nlme)
> data2.gls1 <- gls(y~x, data=data2,
+   weights=varIdent(form=~1|x), method="REML")
```

## 2.9. Heteroscedadacy in ANOVA

```
> library(nlme)
> data2.gls <- gls(y~x, data=data2,
+   method="REML")
> plot(resid(data2.gls,'normalized') ~
+   fitted(data2.gls))
```



```
> library(nlme)
> data2.gls1 <- gls(y~x, data=data2,
+   weights=varIdent(form=~1|x), method="REML")
> plot(resid(data2.gls1,'normalized') ~
+   fitted(data2.gls1))
```





## 2.10. Heteroscedacity in ANOVA

```
> AIC(data2.gls,data2.gls1)
```

```
df      AIC
data2.gls 6 249.4968
data2.gls1 10 199.2087
```

```
> anova(data2.gls,data2.gls1)
```

	Model	df	AIC	BIC	logLik	Test	L.Ratio	p-value
data2.gls		1	6	249.4968	260.3368	-118.74841		
data2.gls1		2	10	199.2087	217.2753	-89.60435	1 vs 2	58.28812 <.0001

- note: it costs d.f.

## 2.11. Heteroscedacity in ANOVA

```
> summary(data2.gls)
```

Generalized least squares fit by REML

Model: y ~ x

Data: data2

AIC	BIC	logLik
249.4968	260.3368	-118.7484

Coefficients:

	Value	Std. Error	t-value	p-value
(Intercept)	40.79322	0.9424249	43.28538	0.0000
xB	5.20216	1.3327901	3.90321	0.0003
xC	13.93944	1.3327901	10.45884	0.0000
xD	-0.73285	1.3327901	-0.54986	0.5851
xE	-10.65908	1.3327901	-7.99757	0.0000

Correlation:

(Intr)	xB	xC	xD	
xB	-0.707			
xC	-0.707	0.500		
xD	-0.707	0.500	0.500	
xE	-0.707	0.500	0.500	0.500

Standardized residuals:

Min	Q1	Med	Q3	Max
-3.30653942	-0.24626108	0.06468242	0.39046918	2.94558782

Residual standard error: 2.980209

Degrees of freedom: 50 total; 45 residual

```
> summary(data2.gls1)
```

Generalized least squares fit by REML

Model: y ~ x

Data: data2



AIC	BIC	logLik
199.2087	217.2753	-89.60435

## Variance function:

Structure: Different standard deviations per stratum

Formula: ~1 | x

## Parameter estimates:

	A	B	C	D	E
	1.00000000	0.91342371	0.40807004	0.08631969	0.12720393

## Coefficients:

	Value	Std.Error	t-value	p-value
(Intercept)	40.79322	1.481066	27.543153	0.0000
xB	5.20216	2.005924	2.593399	0.0128
xC	13.93944	1.599634	8.714142	0.0000
xD	-0.73285	1.486573	-0.492981	0.6244
xE	-10.65908	1.493000	-7.139371	0.0000

## Correlation:

	(Intr)	xB	xC	xD
xB	-0.738			
xC	-0.926	0.684		
xD	-0.996	0.736	0.922	
xE	-0.992	0.732	0.918	0.988

## Standardized residuals:

	Min	Q1	Med	Q3	Max
	-2.3034240	-0.6178652	0.1064904	0.7596732	1.8743230

Residual standard error: 4.683541

Degrees of freedom: 50 total; 45 residual

### 3. Worked Examples

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#### 3.0. Worked Examples